

The maximum estimator is introduced in Problem 7.59 of Devore, *Probability and Statistics for Engineering and the Sciences*, 8th ed., Brooks Cole, 2012. Devore applies it to fake data that is supposed to be waiting times for a bus.

Assume that the pdf is uniform depending on the parameter  $\theta > 0$ ,

$$f(x) = \begin{cases} \frac{1}{\theta}, & \text{if } 0 \leq x \leq \theta; \\ 0, & \text{otherwise.} \end{cases}$$

Integrating gives the cdf

$$F(x) = \begin{cases} 0, & \text{if } x \leq 0; \\ \frac{x}{\theta}, & \text{if } 0 < x \leq \theta; \\ 1, & \text{if } \theta < x. \end{cases}$$

An estimator for  $\theta$  is the maximum of the sample  $X_1, X_2, \dots, X_n \sim \text{Unif}(0, \theta)$ , but as we have seen, it is biased. An unbiased estimator is given by

$$\hat{\theta} = \frac{n+1}{n} \max(X_1, X_2, \dots, X_n).$$

To see this, let us recall the derivation of its pdf. The cdf, by independence

$$\begin{aligned} F(x) &= \mathbf{P}(\hat{\theta} \leq x) \\ &= \mathbf{P}\left(\frac{n+1}{n} \max(X_1, X_2, \dots, X_n) \leq x\right) \\ &= \mathbf{P}\left(\left\{X_1 \leq \frac{nx}{n+1}\right\} \cap \left\{X_2 \leq \frac{nx}{n+1}\right\} \cap \dots \cap \left\{X_n \leq \frac{nx}{n+1}\right\}\right) \\ &= \mathbf{P}\left(X_1 \leq \frac{nx}{n+1}\right) \cdot \mathbf{P}\left(X_2 \leq \frac{nx}{n+1}\right) \dots \mathbf{P}\left(X_n \leq \frac{nx}{n+1}\right) \\ &= F\left(\frac{nx}{n+1}\right)^n \\ &= \begin{cases} 0, & \text{if } x \leq 0; \\ \left(\frac{nx}{(n+1)\theta}\right)^n, & \text{if } 0 < x \leq \frac{n+1}{n}\theta; \\ 1, & \text{if } \frac{n+1}{n}\theta < x. \end{cases} \end{aligned}$$

The pdf for  $\hat{\theta}$  is thus

$$f_{\hat{\theta}}(x) = \frac{d}{dx} F(x) = \begin{cases} \frac{n^{n+1}x^{n-1}}{(n+1)^n\theta^n}, & \text{if } 0 \leq x \leq \frac{n+1}{n}\theta; \\ 0, & \text{otherwise.} \end{cases}$$

Let  $0 < \alpha < 1$ . Let us compute the probability

$$\begin{aligned}
 P\left(\theta\left(\frac{\alpha}{2}\right)^{1/n} \leq \max(X_1, X_2, \dots, X_n) \leq \theta\left(1 - \frac{\alpha}{2}\right)^{1/n}\right) \\
 &= P\left(\theta\left(\frac{\alpha}{2}\right)^{1/n} \leq \frac{n\hat{\theta}}{n+1} \leq \theta\left(1 - \frac{\alpha}{2}\right)^{1/n}\right) \\
 &= F\left(\theta\left(\frac{\alpha}{2}\right)^{1/n}\right) - F\left(\theta\left(1 - \frac{\alpha}{2}\right)^{1/n}\right) \\
 &= \left(\frac{\alpha}{2}\right) - \left(1 - \frac{\alpha}{2}\right) \\
 &= 1 - \alpha.
 \end{aligned}$$

Similarly

$$\begin{aligned}
 P\left(\theta\alpha^{1/n} \leq \frac{n\hat{\theta}}{n+1} = \max(X_1, X_2, \dots, X_n) \leq \theta\right) &= F(\theta) - F(\theta\alpha^{1/n}) \\
 &= 1 - \alpha.
 \end{aligned}$$

These two probability statements become confidence intervals by solving the inequalities for  $\theta$ . With  $1 - \alpha$  confidence, we obtain

$$\frac{\max(X_1, X_2, \dots, X_n)}{\left(1 - \frac{\alpha}{2}\right)^{1/n}} \leq \theta \leq \left(\frac{2}{\alpha}\right)^{1/n} \max(X_1, X_2, \dots, X_n) \quad (1)$$

$$\max(X_1, X_2, \dots, X_n) \leq \theta \leq \frac{\max(X_1, X_2, \dots, X_n)}{\alpha^{1/n}} \quad (2)$$

The gap in inequality (1) exceeds that in (2) because  $2^{1/n} > 1$  and  $2 - \alpha > 1$  imply

$$\frac{1}{\alpha^{1/n}} - 1 < 2^{1/n} \left[ \frac{1}{\alpha^{1/n}} - \frac{1}{(2 - \alpha)^{1/n}} \right] = \left(\frac{2}{\alpha}\right)^{1/n} - \frac{1}{\left(1 - \frac{\alpha}{2}\right)^{1/n}}$$

We plot both widths, the left vs. right sides of the inequality.

We check that the estimator  $\hat{\theta}$  is unbiased. Note that for sample from  $\text{Unif}(0, \theta)$ , the maximum of all  $X_i$  is  $\theta$  so that the statistic ranges in  $0 \leq \hat{\theta} \leq \frac{n+1}{n}\theta$ . Thus

$$\begin{aligned}
 E(\hat{\theta}) &= \int_0^{\frac{n+1}{n}\theta} u f_{\hat{\theta}}(u) du \\
 &= \int_0^{\frac{n+1}{n}\theta} \frac{n^{n+1}u^n}{(n+1)^n\theta^n} du \\
 &= \left[ \frac{n^{n+1}u^{n+1}}{(n+1)^{n+1}\theta^n} \right]_0^{\frac{n+1}{n}\theta} \\
 &= \theta.
 \end{aligned}$$

We simulate  $B = 10,000$  samples of size  $n = 5$  taken from  $\text{Unif}(0, 1)$ . We plot a histogram of the resulting  $\hat{\theta}$ 's and superimpose the density  $f_{\hat{\theta}}$ .

We also simulate  $B = 100$  samples of size  $n = 5$  taken from  $\text{Unif}(0, 1)$  and plot the corresponding  $\alpha = .05$  confidence intervals using (2). It turned out that six of these did not capture the  $\theta = 1$ , which is about the expected 5%.

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**R Session:**

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[Workspace restored from /Users/andrejstreibergs/.RData]

```
> ##### COMPUTE THE TWO CI'S FOR THETA #####
>
> # two ci for th of dist unif(0,th)
> n <- 5
> alpha <- .05
> l1 <- function(y){max(y)*(1-alpha/2)^(-1/n)}
> u1 <- function(y){max(y)*(alpha/2)^(-1/n)}
> l2 <- function(y){max(y)}
> u2 <- function(y){max(y)*(alpha)^(-1/n)}
>
> ##### ENDTER DEVORE'S FAKE TIMES FROM EX. 7.59 #####
> x <- scan()
1: 4.2 3.5 1.7 1.2 2.4
6:
Read 5 items
> # First CI and its width.
> c(l1(x),u1(x));u1(x)-l1(x)
[1] 4.221321 8.783372
[1] 4.562051
> # Second CI and its width.
> c(l2(x),u2(x));u2(x)-l2(x)
[1] 4.20000 7.64637
[1] 3.446370
```

```

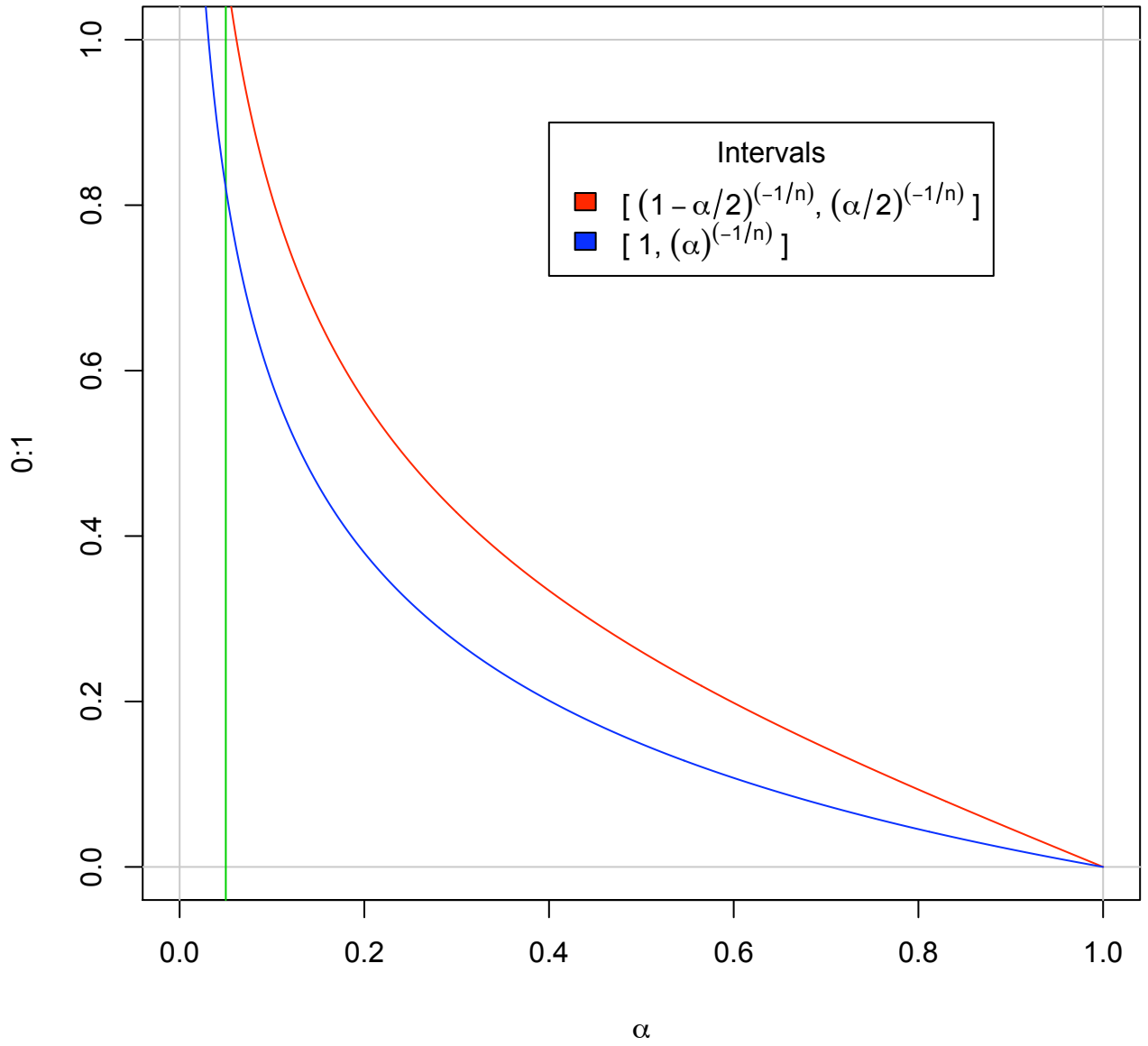
> ##### PLOT TWO WIDTHS TO COMPARE #####
> # Interval width as fn of alpha
> w1 <- function(a){(a/2)^(-1/n)-(1-a/2)^(-1/n)}
> w2 <- function(a){(a)^(-1/n)-1}
> xs <- seq(0,1,1/377)
> plot(0:1, xlim = 0:1, type = "n", ylim = 0:1, main = "Widths of CI's",
+ xlab = expression(alpha))
> abline(h = 0:1, col=8); abline(v = 0:1, col=8); abline(v = alpha, col = 3)
> lines(xs,w1(xs),col=2)
> lines(xs,w2(xs),col=4)
> legend(.4, .9, legend = c(expression(paste("[ ", (1-alpha/2)^(-1/n), " ",
+ (alpha/2)^(-1/n), " ]")), expression(paste("[ 1, ", (alpha)^(-1/n), " ]" ))),
+ fill = c(2,4), bg="white", title = "Intervals")
> # M3074MaxUnifCI1.pdf

> ##### SIMULATE SAMPLING DISTRIBUTION OF THETA HAT #####
>
> B <- 10000
> n <- 5
> c <- (n+1)/n
> v <- replicate(B, c*max(runif(n, 0, 1)))
> hist(v,breaks=40, col=rev(heat.colors(50)))
> hist(v,breaks=40, col=rev(heat.colors(60)))
> summary(v); sd(v)
  Min. 1st Qu.  Median    Mean 3rd Qu.    Max.
0.1720 0.9111  1.0460  1.0010  1.1320  1.2000
[1] 0.1685048

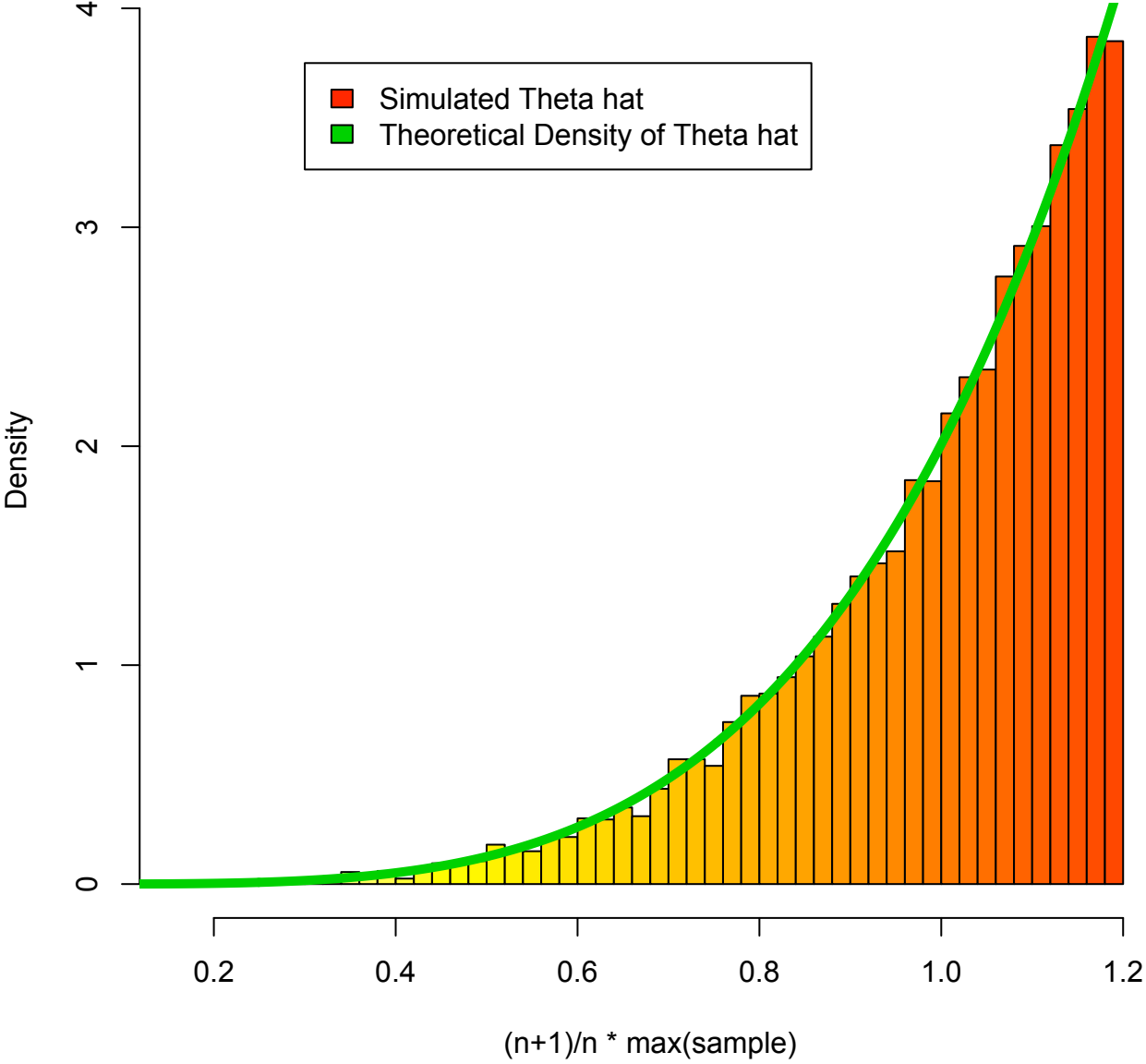
> # Plot the histogram.
> hist(v, breaks = 40, col = rev(heat.colors(60)), xlab="(n+1)/n * max(sample)",
+ main = paste("Simulation of (n+1)/n max, samp.size n =", n,
+ "\n from Unif(0,1), no.reps =", B), freq = F)
> lines(xs, n/c*(xs/c)^(n-1), col = 3, lwd = 5)
> legend(.3, 3.75, legend = c("Simulated Theta hat",
+ "Theoretical Density of Theta hat"), fill = c(2, 3))
> # M3074MaxUnifCI2.pdf

```

### Widths of CI's



**Simulation of  $(n+1)/n$  max, samp.size  $n = 5$   
from  $\text{Unif}(0,1)$ , no.reps = 10000**



```

> ##### SIMULATE 100 CI'S FOR THETA #####
>
> u <- numeric(100); l <- numeric(100)
> # Fill color vector with blues.
> cl <- rep(4,100)
> # Generate CI's. Test if theta is captured.
> for(i in 1:100)
+   {
+     sam <- runif(n,0,1)
+     l[i] <- l2(sam)
+     u[i] <- u2(sam);
+     if(l[i] > 1 | u[i] < 1)
+       {
+         cl[i] <- 2
+       }
+   }
> mx <- max(u)
> mi <- min(l)
> plot(1:100, u, ylim = c(mi, mx), type = "n",
+ main = "Confidence Intervals for Theta")
> for(i in 1:100)
+   {
+     points(c(i, i), c(l[i], u[i]), col = cl[i], pch = 19)
+     lines(c(i, i), c(l[i], u[i]), col = cl[i])
+   }
> abline(h = 1, col = 3)
> legend(65, .5, legend = c("Theta Captured", "Theta Not Captured"),
+ fill = c(4, 2), bg = "white")
> # M3074MaxUnifCI3.pdf

```

### Confidence Intervals for Theta

